

Dispersive evaluations of HVP

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Dispersive evaluations of HVP: outline

- Main objective: evaluation of $a_{\mu}^{\text{HVP,LO}}$ from the $e^+e^- \rightarrow \text{hadrons}$ data base that
 - faithfully propagates the uncertainties from the underlying data
 - optimizes the precision that can be achieved from the combination of data sets

↪ either fully data-driven or including QCD constraints
- Objectives of imposing QCD constraints (analyticity, unitarity, crossing, low-energy theorems)
 - check consistency of data base
 - improve precision
 - establish and explore relations among different channels
- In either case, information on **statistical and systematic uncertainties is crucial**
- Outline for this discussion session:
 - Update on 3π channel MH
 - Update on 2π channel Thomas Leplumey
 - Open discussion

Dispersive evaluations of $e^+e^- \rightarrow 3\pi$

- Data base:

- Data before 2021 [SND](#), [CMD-2](#), . . .
- [BaBar 2021](#), [2110.00520](#)
- [Belle II 2024](#), [2404.04915](#)
- [SND 2026](#), [2603.01635](#)

- Dispersive representation based on:

- Khuri–Treiman equations for $\gamma^* \rightarrow 3\pi$
- Normalization function $a(q^2)$ in terms of $\omega, \phi, \omega', \omega''$ resonances and conformal polynomial
- Low-energy theorem for $\gamma\pi \rightarrow \pi\pi$ (chiral anomaly C_A , fixed in default fit)
- ρ – ω mixing via coupled-channel formalism

- Errors will be given as:

- 1 Experiment (including scale factor $S = \sqrt{\chi^2/\text{dof}}$)
- 2 Truncation of conformal expansion
- 3 Phase of ρ – ω mixing parameter
- 4 Sum in quadrature of 1.–3.

Dispersive evaluations of $e^+e^- \rightarrow 3\pi$

	Belle II	BaBar	SND	Global fit (BaBar+SND+earlier) (w/o Belle II)
χ^2/dof	$\frac{218.2}{159} = 1.77$	$\frac{183.7}{129} = 1.42$	$\frac{134.7}{126} = 1.07$	$\frac{672.8}{467} = 1.44$
M_ω [MeV]	782.72(5)(0)(0)[5]	782.542(11)(1)(0)[11]	782.574(10)(4)(2)[11]	782.585(10)(1)(2)[11]
Γ_ω [MeV]	8.43(8)(1)(2)[8]	8.718(26)(2)(3)[26]	8.631(19)(14)(7)[25]	8.672(15)(12)(9)[21]
M_ϕ [MeV]	1019.24(5)(1)(2)[5]	1019.276(6)(0)(0)[6]	1019.258(14)(25)(4)[29]	1019.246(11)(7)(4)[14]
Γ_ϕ [MeV]	3.98(11)(2)(0)[11]	4.289(10)(0)(0)[10]	4.271(26)(16)(4)[31]	4.279(9)(1)(1)[9]
$a_\mu^{3\pi} _{\leq 1.8 \text{ GeV}}$	48.7(1.4)(0.1)(0.1)[1.4]	45.99(53)(15)(6)[55]	45.91(35)(15)(2)[38]	45.91(28)(16)(9)[33]
C_A/C_A^{WZW}	1.069(58)(29)(12)[66]	0.945(27)(14)(18)[36]	1.005(19)(51)(16)[57]	1.024(14)(38)(16)[44]

● Belle II data do not pass the test

- Bad fit quality
- Resonance parameters off
- (Strong) pull in C_A , see below

↪ enough cause to discard?

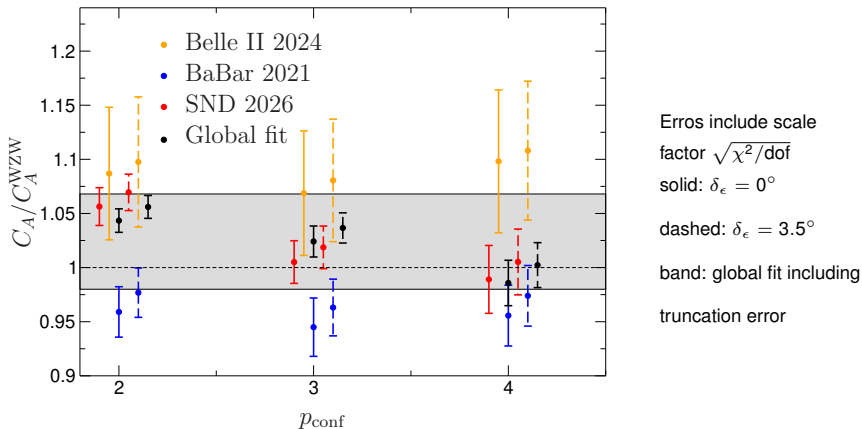
Dispersive evaluations of $e^+e^- \rightarrow 3\pi$

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- BaBar and SND look amazingly consistent
- Treatment of systematic errors for SND unclear
 - L , $N_{3\pi}$, ϵ , RC each fully correlated or not?
 - If yes, then fit explodes
 - Above results based on best guess for covariance matrix by Aidan Wright from 2603.01635

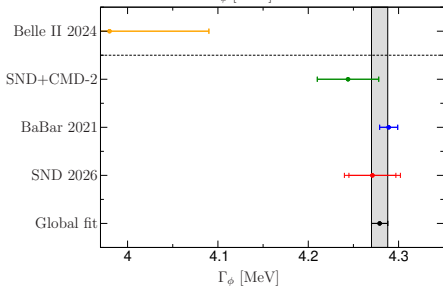
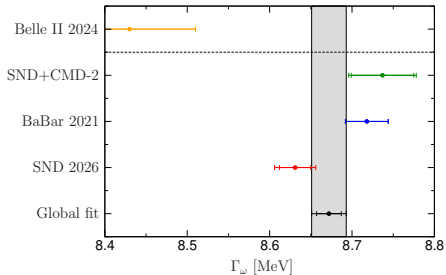
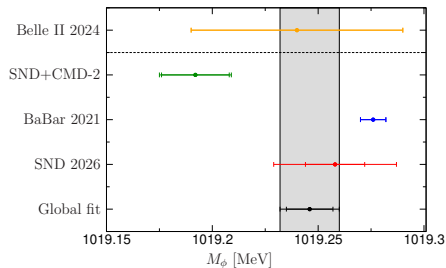
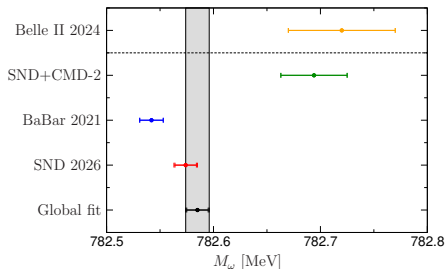
↪ should we really guess the correlations ourselves?

Dispersive evaluations of $e^+e^- \rightarrow 3\pi$: chiral anomaly



- Main systematic error from truncation of conformal polynomial p_{conf}
- BaBar, SND, and global fit converge towards WZW expectation
- Belle II systematically above

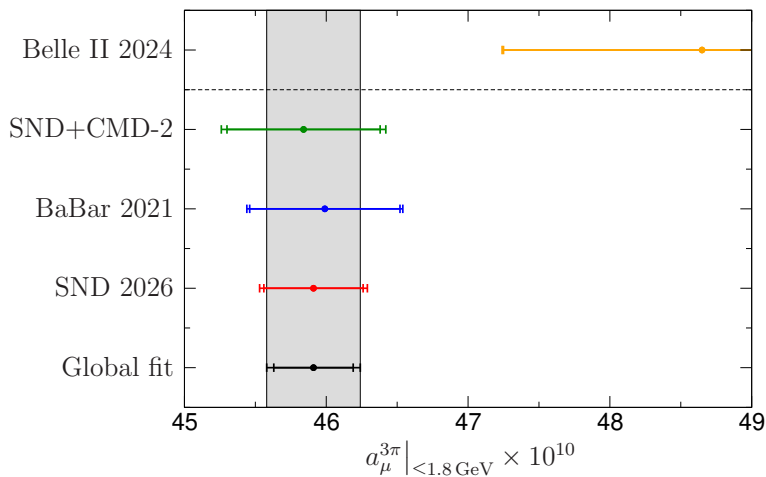
Dispersive evaluations of $e^+e^- \rightarrow 3\pi$: resonance parameters



All parameters exclude vacuum polarization

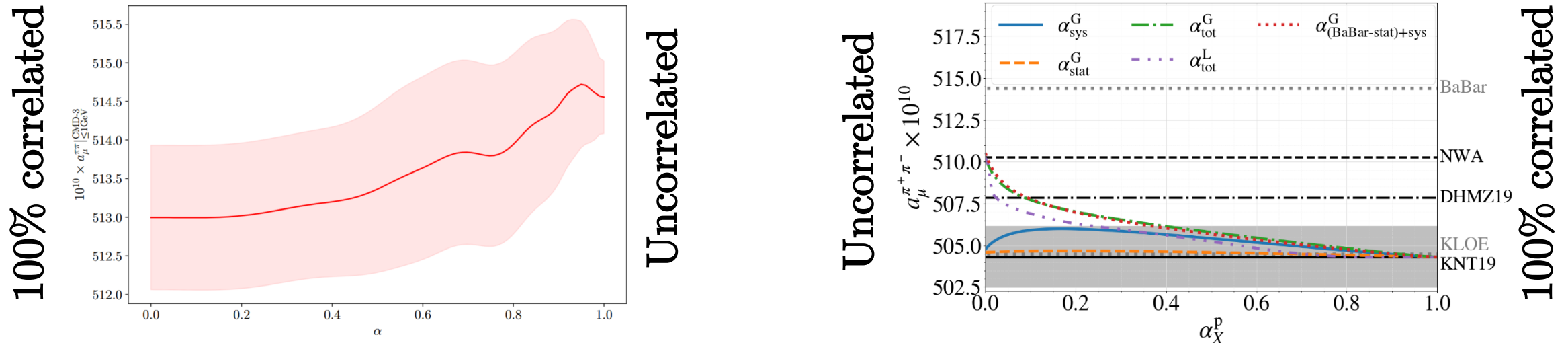
BaBar errors for $M_{\omega,\phi}$ do not account for energy-calibration uncertainty

Dispersive evaluations of $e^+e^- \rightarrow 3\pi$: contribution to a_μ



Uncertainties on uncertainties

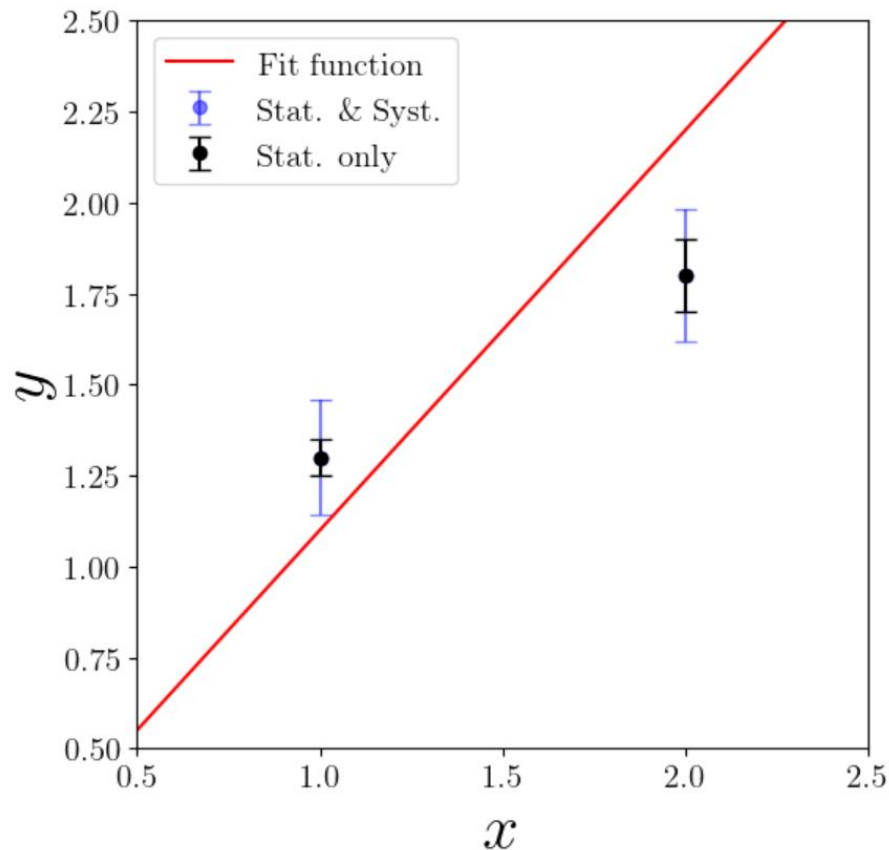
- Nominal CHKLS fit results previously assumed 100% correlated syst. when systematic covariance was not provided by the experiments
- But we are now aware that such an assumption **influences a lot the results**
 - arXiv:2501.09643 (Leplumey, Stoffer): implements a simple local decorrelation scheme
 - arXiv:2604.25004 (KNTW): more detailed studies, interplay local & global decorrelation schemes



1. We would like to find a procedure that **faithfully covers** the post-fit constraints given by **all possible correlations** at realistic computation costs
2. To what extent do we need to consider « all possible correlations »? Can we define the coverage with respect to a **simpler / more relevant ensemble of correlations**?

Toy model

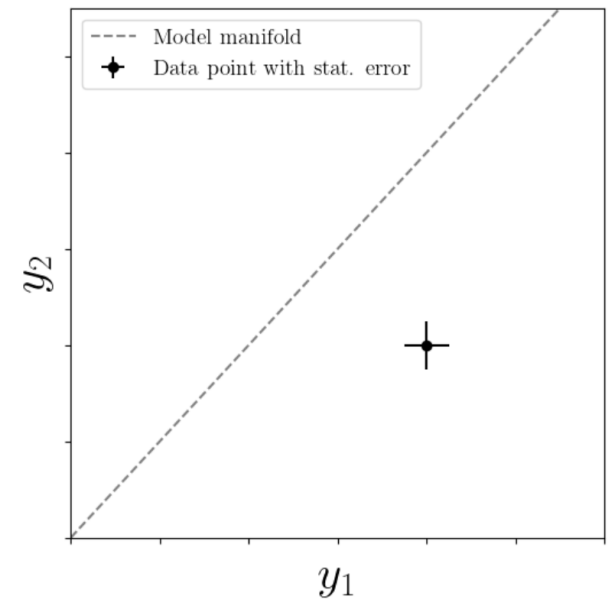
- Consider a toy model:
 - 2 data points, diagonal statistical covariance
 - Fit with a linear function $y = ax$
 - Given systematic uncertainty for each data point, but unknown correlation



2D Observable space: (y_1, y_2)

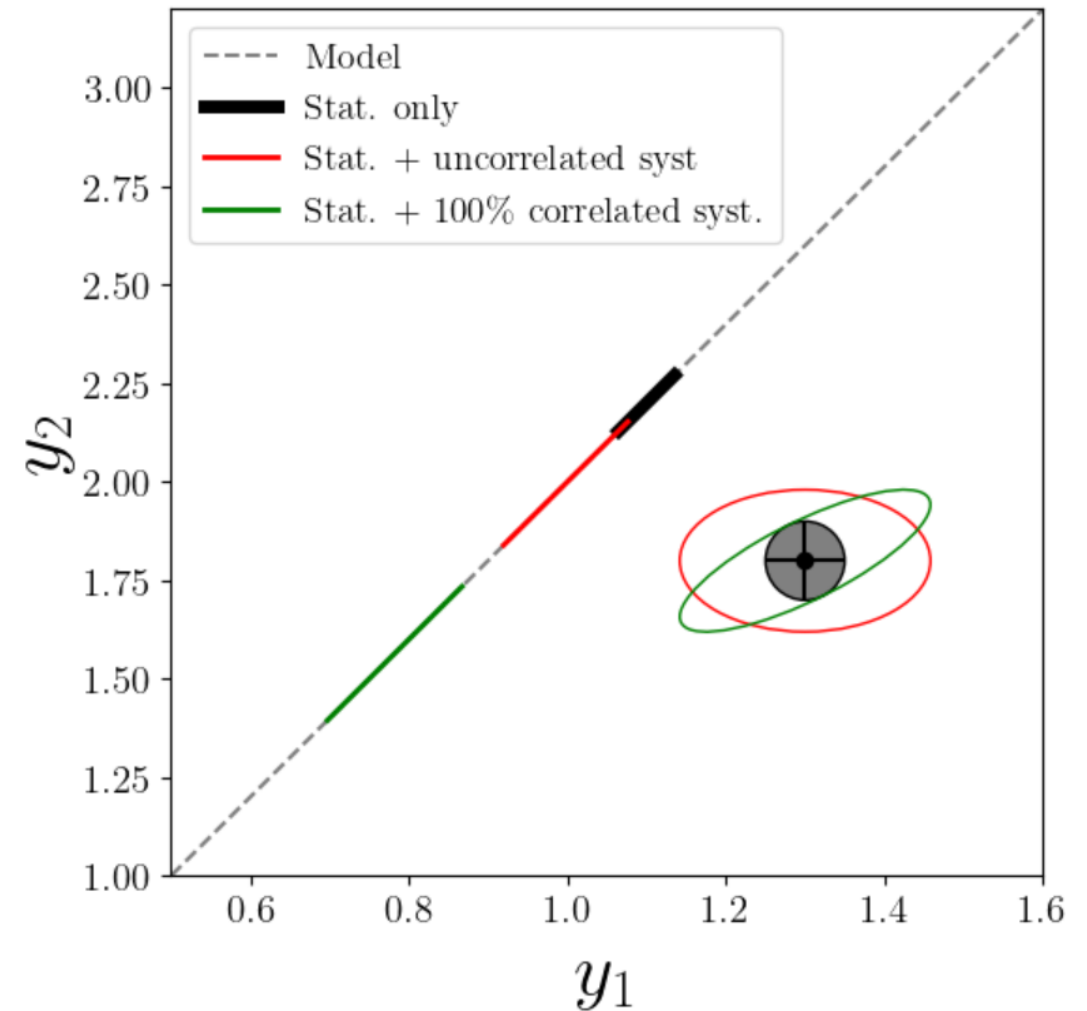
1D Model manifold: $\{(ax_1, ax_2), a \in \mathbb{R}\}$

We can represent the observed data point and the model manifold in this 2D observable space



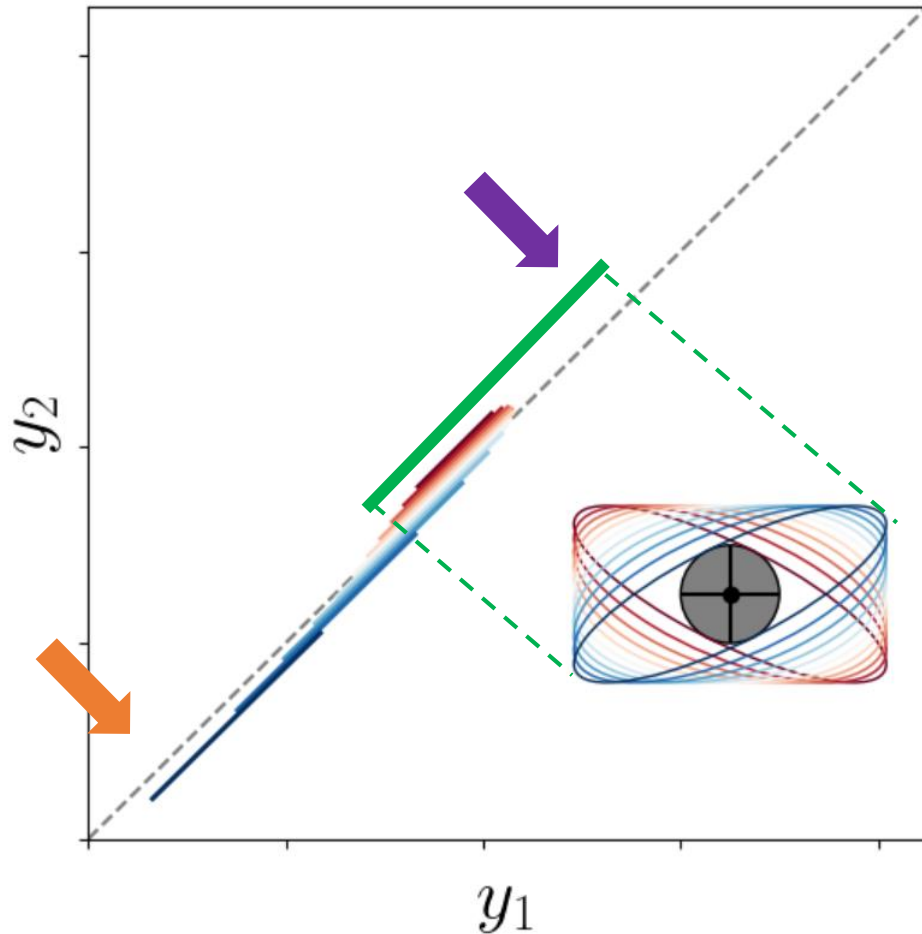
Impact of different correlation assumptions on toy model

- This toy model illustrates the effect of **changing the syst. correlation factor**
 - Size of the post-fit uncertainty
 - Shift in the best-fit parameters
 - Goodness of fit
 - Fit $\approx$ projection of data on model manifold wrt given covariance matrix
 - Distance to model manifold = goodness of fit
 - Stat only: very bad GoF
 - **Uncorrelated syst: good GoF**
 - **100% correlated: poorer GoF**
- } *Big shift*



Effect of the method proposed at last workshop

- By varying the correlation factor between -1 and 1 , we can determine the range spanned within the model manifold



- In reality: $O(10^2)$ data points i.e. $O(10^4)$ correlation factors
→ we can't just scan all of them
- Proposal at last Workshop: Shift the data points one by one within their syst errors and add the shifts linearly
→ **result shown on toy model**
- Not representative of the fit (disregards model constraints)
 - **Over-coverage**: covered regions with very bad GoF
 - **Under-coverage**: uncovered regions favored by the model

How would we want to treat this problem?

- Mathematical reason why this procedure does not do what we want:

$$\Sigma_{\text{params}} = \left[P^T (\Sigma_{\text{data}}^{\text{stat}} + \Sigma_{\text{data}}^{\text{syst}})^{-1} P \right]^{-1} \quad \text{Non-linear!} \quad \text{(could work if } \Sigma_{\text{syst}} \text{ was small compared to } \Sigma_{\text{stat}}, \text{ but it's not the case)}$$

- In general, finding a robust bound that is not over-conservative and without making any assumption about the correlations is non-trivial
 - Attempts using e.g. Woodbury identity, but not converged yet

$$(A + UCV)^{-1} = A^{-1} - A^{-1}U(C^{-1} + VA^{-1}U)^{-1}VA^{-1},$$

- A more realistic approach is to scan only a **sub-class of correlations matrices**
 - KNTW approach in arXiv:2604.25004: parameterized “global” and “local” decorrelations
 - We can **extend this to other or larger classes**, if we have good motivations
- But when/how do we consider that we have a **representative coverage**?

Another practical attempt

- Early attempts we made tried to find the correlations that maximize the post-fit uncertainty

- Assume some correlations as a starting point
- Using first order approx, find the corr matrix that maximize $V[a_\mu^{\pi\pi}]$
- Re-fit using this new correlation matrix, and iterate until it's stable

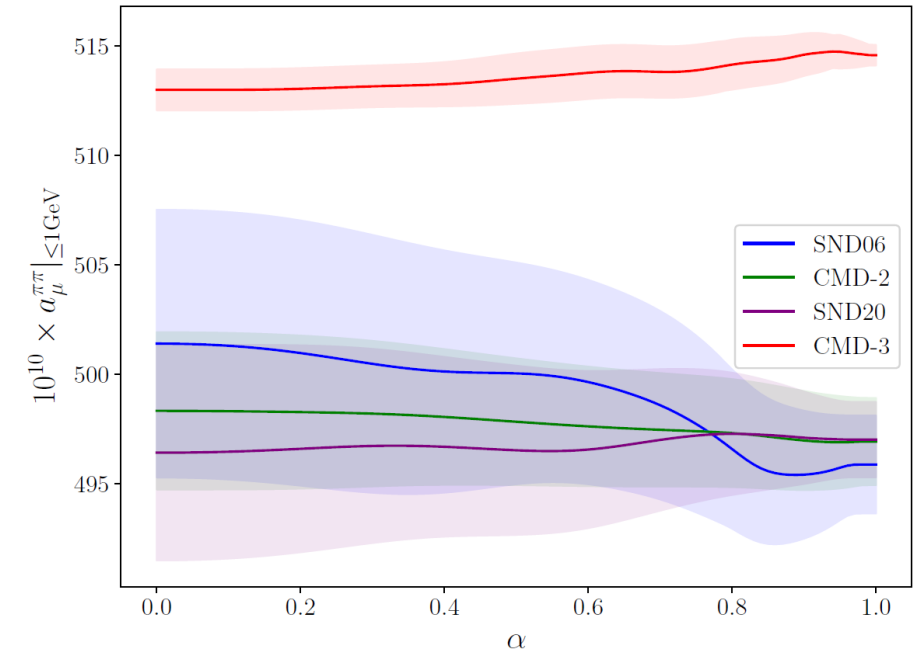
$$\text{Corr}(\sigma_i, \sigma_j) = \text{sign} \left(\frac{\partial a_\mu^{\pi\pi}}{\partial \sigma_i} \times \frac{\partial a_\mu^{\pi\pi}}{\partial \sigma_j} \right)$$

- But we observed that the result changes when we change the starting point
 - This stresses out that **1st order approximation fails for this use case** (non-linear manifold)
 - Furthermore, we now understand that this is not a good approach due to best-fit shift
- Follow-up, to be checked: find correlations that **min(maximize the value of $a_\mu^{\pi\pi}$ and take full interval as syst. uncertainty** (incorporates shifting effects)
 - But no guarantee yet that we wouldn't have the same issue...

- On top of giving **different model constraints**, the different assumptions on correlations give **extremely different goodness of fit** (sometimes absurd)
 - e.g. $\frac{\chi^2}{\text{dof}} \sim 0.5$ in CMD-3 uncorrelated case, while $\frac{\chi^2}{\text{dof}} \sim 1.1$ for 100% correlated case...
- Thus, defining 1σ region as the envelope of all possible cov. matrices would not work (only the edge with best GoF would matter)
 - Difficult to find a prescription inherently **compatible with frequentist approach**...
- Also an issue to determine the **compatibility of different datasets** with unitarity, analyticity constraints, etc.
 - Would it be enough if one set of correlations gives a good GoF to say it's compatible?

Our conclusions so far

- Empirically, uncorrelated and 100% correlated syst. seem to provide **nearly maximal coverage**, though we cannot provide a solid proof of this
- For CMD-3, decorrelation pulls $a_{\mu}^{\pi\pi}$ upward, while for most other experiments it pulls $a_{\mu}^{\pi\pi}$ downward
→ Claiming discrepancies with 100% correlations is somewhat conservative



Our wish list

- Nevertheless, **guessing what the correlations should be cannot be fully a theorist task**
- Experiments should **provide all the information they have** as much as possible
- For instance, releasing several cov. matrices **for each systematic source**, and **provide correlations for those which are available** (e.g. detector effects)

- How can we improve the situation regarding systematic uncertainties?
 - Models for decorrelation
 - Casting information into the form of covariance matrices
- Interplay between systematic and statistical correlations
- Under which conditions can we discard experiments in global compilations?
SND 20 for 2π , Belle II for 3π
- Can the evidence derived from inelastic channels be made more significant?
- ...